The Role of CDS Trading in the Commercialization of New Lending Relationships

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Abstract

We investigate how the development of the credit default swap (CDS) market affects lenders' incentives to initiate new lending relationships. We predict that hedging loan exposure through CDSs and the revelation of private information through CDS spreads reduce the adverse selection that non-relationship lenders face when competing for loans. Consistent with our expectations, we find that, following CDS initiation on a borrower's debt, non-relationship lead arrangers are more likely to originate its loans and non-relationship participants are more likely to join loan syndicates. We also show that, following CDS initiation, non-relationship lenders pursue new borrowers with high cross-selling potential and that lenders with low loan growth or profitability are more likely to initiate new relationships. We further find that, relative to relationship lead arrangers, non-relationship lead arrangers have lower monitoring incentives following CDS initiation, as reflected by their weaker control rights and lower loan share retained. As these lower monitoring incentives exacerbate agency problems within the syndicate, the increase in the likelihood of a non-relationship participant joining the syndicate following CDS initiation is attenuated if a loan is syndicated by a non-relationship lead arranger. Overall, our findings shed light on how CDS trading affects the evolution of lending relationships.

1. Introduction

The development of the credit default swap (CDS) market is one of the most important financial innovations of recent decades. CDSs offer lenders the chance to hedge borrowers' credit risk while maintaining their lending relationships (e.g., Saretto and Tookes 2013). However, hedging through CDSs reduces lenders' incentives to monitor borrowers and may push borrowers into inefficient bankruptcy (e.g., Hu and Black 2008; Ashcraft and Santos 2009; Bolton and Oehmke 2011; Parlour and Winton 2013). Considering these positive and negative attributes of CDS trading, we examine how the availability of CDSs influences the loan origination and participation decisions in loan syndicates.

We predict that non-relationship lenders (i.e., lenders without a relationship with a borrower before CDS initiation) are more likely to serve as lead arrangers or join loan syndicates as participants once CDSs on a borrower's debt become available. Non-relationship lenders competing for a loan face adverse selection risk, as relationship lenders have a substantial information advantage over them, due to prior experience with the borrower (Boot 2000). Because relationship lenders exploit this advantage and bid for good loans while avoiding bad ones, non-relationship lenders will compete less aggressively for a loan, which typically results in them losing a loan to a better-informed incumbent lender (Rajan 1992). But once CDSs become available, an opportunity to hedge against potential credit losses should significantly mitigate non-relationship lenders' adverse selection concerns. CDS spreads also often reveal private information about a borrower, diminishing relationship lenders' information advantage (e.g., Acharya and Johnson 2007; Qiu and Yu 2012; Batta et al. 2016). We expect lower adverse selection to increase the willingness of non-relationship lenders to bid for a loan, thus increasing their probability of winning a deal. This reasoning applies to non-relationship lenders competing

with incumbent lead arrangers to serve as the lead arranger of a borrower's loan as well as to non-relationship lenders competing to join the syndicate with other participants that have superior information about the borrower through prior lending.

Using staggered initiation of CDS trading across CDS firms, we employ a difference-in-differences research design with firm and time fixed effects to explore our predictions. We find a positive and significant effect of CDSs on non-relationship lending for both lead arrangers and syndicate participants. The probability that a non-relationship lead arranger (syndicate participant) arranges (joins) the loan syndicate is 1.39 (1.28) times higher following CDS trading initiation on a borrower's debt. These findings are robust when we perform short-window analyses around CDS initiation, propensity-score match CDS firms with non-CDS firms, and employ the instrumental variable approach. Overall, our analyses suggest that, due to CDS hedging and the revelation of private information through CDS spreads, lenders are more likely to initiate new lending relationships once CDSs become available.

To better understand the underlying mechanisms of non-relationship lending decisions following CDS initiation, we next explore whether non-relationship lenders are more likely to pursue borrowers with high cross-selling potential. Lending decisions are typically driven by the overall profitability of the relationship with the borrower, including non-interest income (Standard & Poor's 2011). Lenders cross-sell various services, such as investment banking, derivatives, and structured finance, generally aiming to serve as a "one-stop shop" of all financial services for their borrowers (Fang et al. 2013). We predict that, because non-relationship lenders are likely to hedge their loan exposure, thus reducing concerns regarding a borrower's credit risk, they focus more on commercial aspects of the lending relationship and pursue new borrowers with high cross-selling potential. Substantial fees generated by cross-selling additional services

to the borrower also help compensate for the cost of purchasing CDS protection, further reinforcing our prediction.

Because lenders' expectations about borrowers' future business needs are difficult to assess, we use multiple measures of cross-selling potential: 1) media coverage of a borrower's activities, including public offerings, mergers and acquisitions (M&A), and investments; 2) a borrower's M&A after a loan's issuance; 3) whether a borrower is in the aviation, ship building, oil, refinery, gas, or telecommunication industries, which offer lenders opportunities to sell structured finance products; 4) a borrower's use of derivatives; and 5) the extent of a borrower's foreign operations. For all measures, we find that, following CDS initiation, lead arrangers are more likely to enter into new lending relationships if a borrower has high cross-selling potential. The likelihood of non-relationship participants joining the syndicate is also enhanced by cross-selling potential, but to a lesser degree, which is likely explained by lead arrangers being the primary beneficiaries of cross-selling opportunities (e.g., Ivashina and Kovner 2011).

To shed additional light on lenders' incentives to initiate new lending relationships following CDS initiation, we examine whether CDSs have a stronger effect on non-relationship lenders that experience low loan growth, likely due to internal credit risk-management restrictions or regulatory capital constraints. Since CDSs offer more flexible credit risk management and alleviate capital constraints, we predict that these lenders are more likely to seek new borrowers after the onset of CDS trading. We also expect lenders with low profitability to have stronger incentives to initiate new relationships once they can hedge their exposure through CDSs, as these relationships may boost profitability by generating additional interest revenue and cross-selling fees. Consistent with these predictions, we find that low loan growth and profitability amplify the effect of CDSs on the probability of initiating new relationships for

lead arrangers and syndicate participants.

We extend our analyses by testing non-relationship lenders' monitoring incentives. These analyses are based on lead arrangers' relationship with the borrower, as lead arrangers perform the primary monitoring of the borrower on behalf of syndicate participants. Although prior studies show that CDSs diminish the strength of lead arranger monitoring (e.g., Subrahmanyam et al. 2014; Martin and Roychowdhury 2015), we expect this adverse effect to be stronger for non-relationship lead arrangers. Because non-relationship arrangers start syndicating loans to a borrower only once CDSs become available, they are more likely to rely on CDS hedging, rather than intensive monitoring, to protect themselves against credit risk. Weak monitoring may also help non-relationship arrangers attract new borrowers, as well as give these borrowers flexibility to pursue new business activities, which in turn creates further cross-selling opportunities.

We measure lenders' monitoring incentives by the strength of their control rights, since lenders retain stronger rights when they intend to monitor a borrower more closely (Roberts and Sufi 2009a, 2009b; Roberts 2015). Loan contracts with a higher ratio of performance (income statement-based) covenants to total financial covenants and those that incorporate interest-increasing performance-pricing provisions give lenders stronger control rights (Roberts and Sufi 2009a; Christensen and Nikolaev 2012; Christensen et al. 2016). Consistent with our prediction, we find that, following CDS initiation, non-relationship lead arrangers impose fewer performance covenants relative to total financial covenants and are less likely to impose interest-increasing provisions relative to relationship lead arrangers. Note that we find that lead arrangers that had a relationship with the borrower before CDS initiation do not relinquish their control rights afterward. This evidence indicates that the decrease in lenders' control rights following CDS initiation, documented previously (e.g., Shan et al. 2015; Chakraborty et al. 2015), stems

primarily from lenders that initiate a relationship with the borrower after CDS initiation.

Lead arrangers also have stronger monitoring incentives when they have more skin in the game, i.e., retain a larger loan share (e.g., Sufi 2007; Ivashina 2009). Therefore, we predict that, following CDS initiation, non-relationship lead arrangers retain a lower loan share relative to relationship lead arrangers. We find strong support for this prediction. This evidence also extends Amiram et al.'s (2017) findings of lead arrangers' larger skin in the game following CDS initiation, as we show that these findings are explained primarily by lenders who had relationships with borrowers before CDS initiation.

Building on the stronger negative influence of CDSs on monitoring of non-relationship relative to relationship lead arrangers, we expect syndicate participants to be concerned that the former will shirk more on their monitoring duties. The agency problems within a syndicate are mitigated if lenders with prior relationships with the borrower join the syndicate, since they are more informed about the borrower's performance and creditworthiness (e.g., Sufi 2007; Ivashina 2008). Thus, although we hypothesize and find that non-relationship participants are more likely to join the syndicate following CDS initiation, we predict that this effect will be attenuated if a non-relationship lead arranger syndicates a loan. We indeed find that, following CDS initiation, the probability that a non-relationship participant joins a syndicate is lower if the loan is arranged by a non-relationship relative to relationship lead arranger.

We conduct additional exploratory tests of the effect of CDSs on loan terms. To disrupt existing lending relationships, non-relationship lenders may entice borrowers with favorable loan terms, such as a low interest rate spread and large loan size. However, because hedging via CDSs is costly, these terms may diminish the benefits of the new lending relationship. We find that, following CDS initiation, non-relationship lead arrangers do not charge lower spreads but do

issue smaller loans, relative to relationship lead arrangers. Coupled with the significant impact of cross-selling potential on non-relationship lenders' decisions, this evidence suggests that, even though these lenders provide new borrowers with less credit to reduce hedging costs, they still get a mandate to provide them with additional financial services.

Our paper contributes to the literature on the influence of the CDS market on private lending. Prior studies find a significant impact of CDSs on both borrowers' access to credit and the strength of lenders' monitoring (e.g., Ashcraft and Santos 2009; Saretto and Tookes 2013; Subrahmanyam et al. 2014; Martin and Roychowdhury 2015). These studies, however, mostly do not differentiate between existing and new lending relationships. We extend this literature by exploring how CDSs affect the formation of new relationships and the multifaceted lending strategies of non-relationship lenders. We show that CDS trading increases the likelihood of initiating new relationships for both lead arrangers and syndicate participants, especially when a borrower has high cross-selling potential or lenders experience low loan growth and profitability. Further, we supplement prior studies that examine the underlying mechanisms through which CDSs affect lenders' monitoring incentives, such as the loan share retained by the lead arranger

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¹ In a concurrent paper, Shan et al. (2016) also document higher likelihood of loans syndicated by non-relationship lead arrangers following CDS initiation. They attribute this finding to borrowers switching to new lead arrangers because borrower-lead arranger relationship is compromised by CDSs and becomes less valuable to the borrower. In contrast, building on Rajan (1992), we suggest that CDSs decrease adverse selection between incumbent and nonrelationship lenders, thus increasing the willingness of the latter to compete for new borrowers and consequently the likelihood of their winning a loan deal. Shan et al.'s (2016) proposition of compromised lead arranger-borrower relationships cannot explain the majority of our findings, including 1) the effect of CDSs on the probability of nonrelationship participants joining syndicates; 2) the amplifying effect of lenders' low loan growth and profitability on initiating new lending relationships; and 3) the importance of a borrower's cross-selling potential for lenders' decisions to enter new lending relationships. Note, too, that our findings with respect to lead arrangers' monitoring incentives further undermine Shan et al.'s (2016) motivation. They argue that borrowers switch away from relationship lenders primarily because they are concerned about these lenders' monitoring incentives. In contrast, we show that relationship lenders actually do not relinquish their control rights and increase their "skin in the game" following CDS initiation, while non-relationship lenders monitor substantially less intensively than relationship ones. In addition, our findings with respect to the interest spread differ from those of Shan et al. (2016), potentially due to the fact that they do not employ a stringent difference-in-differences research design.

and lead arrangers' control rights (e.g., Chakraborty et al. 2015; Shan et al. 2015; Amiram et al. 2017). We show that the influence of CDS trading on these mechanisms depends to a large extent on whether lenders had an established relationship with a borrower before CDS initiation.

We also contribute to the large literature on relationship lending. While some studies show important benefits to borrowers from established lending relationships (e.g., Petersen and Rajan 1994, 1995; Berger and Udell 1995; Brahrath et al. 2009), others emphasize its dark side, where borrowers become locked in a relationship, due to the information advantage of incumbent lenders (e.g., Sharpe 1990; Rajan 1992). We provide evidence consistent with CDS trading fundamentally altering competition in the loan market by enhancing the willingness of non-relationship lenders to compete for new borrowers. Relatedly, our findings underscore important implications of financial innovation on the evolution of the lender-borrower relationship.

Finally, we add to the growing literature on the importance of noncredit revenue sources in private lending. Prior studies show that lenders gain future underwriting of a borrower's bonds and equity (e.g., Yasuda 2005; Drucker and Puri 2005). There is also evidence that lenders price in the loan spread a borrower's cross-selling potential as well as its board network, as this network offers an advantage in cross-selling services to other firms (e.g., Ivashina and Kovner 2011; Zhao 2017). We complement these studies by showing that a borrower's cross-selling potential is instrumental in lenders' decisions to initiate new relationships.

The next section presents hypotheses development. Section 3 describes data and sample selection. Section 4 reports our findings, and Section 5 concludes.

2. Related Literature and Hypothesis Development

2.1 CDS trading and non-relationship lending in the syndicated loan market

The introduction of CDSs significantly influenced the private debt market by offering

banks hedging opportunities to lay off borrowers' credit risk, while maintaining lending relationships (e.g., Saretto and Tookes 2013). Even if lenders do not hedge their loan exposure at loan initiation, the existence of CDSs offers them a liquid resale option and thus increases their willingness to extend credit. Further, CDSs allow lenders more flexible risk management and provide them an opportunity to reduce regulatory capital requirements by substituting the risk weight of the CDS counterparty (typically, a large financial institution) for that of the borrower (Saretto and Tookes 2013; Streitz 2015; Martin and Roychowdhury 2015; Shan et al. 2016).

Yet CDSs also induce empty creditor problems as CDS-protected lenders become intransigent in debt renegotiation or even push borrowers into inefficient liquidation (Hu and Black 2008; Bolton and Oehmke 2011). CDSs reduce lenders' monitoring incentives because hedging their loan exposure may be less costly alternative to protect themselves against default risk than intensive information collection and monitoring. Consistent with weaker monitoring, prior studies find that, following CDS initiation, the cost of debt increases for risky and informationally opaque borrowers, borrowers experience more credit rating downgrades and bankruptcies and report less conservatively, and lenders impose looser financial covenants and weaker restrictions after covenant violations (Ashcraft and Santos 2009; Subrahmanyam et al. 2014; Martin and Roychowdhury 2015; Chakraborty et al. 2015; Shan et al. 2015). Amiram et al. (2017) show that syndicate participants, who largely delegate monitoring to the lead arranger, require the arranger to have more skin in the game to compensate for the reduction in its monitoring incentives after the initiation of CDS trading.

We extend these studies by examining how the availability of CDSs affects the likelihood of lenders without a prior relationship with the borrower serving as the lead arranger or joining its loan syndicate as a participant. Because relationship lenders have an information advantage,

due to their extensive knowledge of a borrower's operations and creditworthiness, non-relationship lenders competing for a borrowers' loans face substantial adverse selection risk (Boot 2000). This typically results in a non-relationship lender losing competition for a borrower's loan to a better-informed incumbent lender. Specifically, Rajan (1992) shows that relationship lenders exploit their information advantage and bid for good loans while avoiding bad ones. Because relationship lenders avoid bad loans, non-relationship lenders face significant losses if they do bid. As a result, non-relationship lenders, being aware of the adverse selection problem, compete less aggressively, which reduces the probability that they win the loan.

The availability of CDSs on a borrower's debt mitigates adverse selection risk that non-relationship lenders face because it allows them to hedge loan exposure, thus assuaging their concerns regarding potential credit losses when initiating a new relationship. CDS spreads also often reveal private information about a borrower, ahead of public disclosures and price discovery in other markets, largely due to insider trading of informed financial institutions (e.g., The Financial Times 2005; Acharya and Johnson 2007; Qiu and Yu, 2013; Batta 2016). The revelation of a substantial private information through CDSs should reduce relationship lenders' information advantage, further alleviating non-relationship lenders' adverse selection concerns.

This reasoning applies to lead arrangers as well as syndicate participants. Relationship lenders and non-relationship lenders compete to serve as the lead arranger on a borrower's loan. Lower adverse selection, due to CDS availability, is expected to increase non-relationship lead arrangers' willingness to compete for a borrower's loan, increasing the probability of their winning the deal (Rajan 1992). Similarly, lenders compete to join syndicates. Syndication represents an auction in which loan participants submit sealed bids to the lead arranger, and the number of invitations extended to potential participants typically exceeds the number of lenders

who will fund the loan (Champagne and Kryzanowski 2007; Ivashina and Sun 2011). Lenders with superior information about the borrower, because of prior lending relationships, are likely to pursue participation in good loans, leaving bad ones to non-relationship participants. We expect lower adverse selection, due to CDS hedging and the revelation of private information through CDS spreads, to increase the willingness of non-relationship lenders to bid for loan participation. Building on these argument, we state our first hypothesis as follows.

H1: The probability that a non-relationship lead arranger (syndicate participant) arranges (joins) the loan syndicate is higher following the initiation of CDS trading on a borrower's debt.

2.2 The Importance of Cross-selling Opportunities in Non-Relationship Lending

We next investigate whether a borrower's cross-selling potential intensifies the effect of CDS trading on non-relationship lending in the syndicated loan market. Lenders invest in loans for more than just the interest income: they consider the overall profitability of the relationship, including noncredit revenue (Standard & Poor's 2011). In particular, lenders typically adopt a portfolio-management approach, in which they allocate capital to borrowers based on the total return generated by a relationship, relative to its risk.

Lenders cross-sell various services, including investment banking, derivatives, and structured finance. Yasuda (2005) and Drucker and Puri (2005) show that serving as a lead arranger helps banks to gain future underwriting of a borrower's bonds and equity. Ivashina and Kovner (2011) also show that lenders account for cross-selling potential when pricing leveraged buyout loans. Fang et al. (2013) provide evidence consistent with banks aiming to serve as "one-stop shop" for financial services for their borrowers.

Because non-relationship lenders initiate lending relationship only ones CDSs become available, they are likely to hedge their exposure, thus diminishing concerns about a borrower's

credit risk. Thus, we expect them to focus more on commercial aspects of the lending relationship and pursue new borrowers with high cross-selling potential. Further, cross-sold products are mostly fee based and entail minimal risk-weighted assets. Therefore, fees from cross-selling help lenders compensate for the cost of purchasing CDS protection, while the services sold impose only insignificant pressure on regulatory capital ratios. We state our second hypothesis as follows.

H2: The increase in the probability that a non-relationship lead arranger (syndicate participant) arranges (joins) the loan syndicate following the initiation of CDS trading on a borrower's debt is higher when the borrower has a high cross-selling potential.

2.3 Lender Characteristics and Non-Relationship Lending

To illuminate lenders' incentives to initiate new relationships with borrowers after the onset of CDS trading, we next investigate how lender characteristics affect the relation between CDS trading and non-relationship lending. Lenders often cannot expand their loan portfolio, because they are restricted by an internal credit risk management policy or face regulatory capital constraints. CDS hedging offers them more flexible credit risk management strategies and alleviate regulatory constraints (Streitz 2015; Shan et al. 2016). Therefore, we expect that lenders experiencing low loan growth will be more likely to extend their loan portfolio and seek new borrowers following CDS initiation.

We also expect CDS initiation to more strongly affect less profitable lenders. When hedging via CDSs and information revelation through CDS spreads become available, these lenders are likely to be eager to pursue new relationships to enhance profitability. By expanding loan portfolios, they will benefit from higher interest income and fees from cross-selling additional services to new borrowers. We state our third hypothesis as follows.

H3: The increase in the probability that a non-relationship lead arranger (syndicate participant) arranges (joins) the loan syndicate following the initiation of CDS trading on a borrower's debt is higher for arrangers (participants) with low loan growth or profitability.

2.4 Non-relationship Lenders' Monitoring Incentives

We further study how incentives and strategies of CDS-protected non-relationship lenders affect their monitoring incentives. Here, we focus on lead arrangers of syndicates, because they negotiate loan terms with the borrower and assume the primary screening and monitoring responsibilities on behalf of syndicate participants (e.g., Lee and Mullineaux 2004; Sufi 2007; Ivashina 2009). Even though prior studies document that CDSs undercut lenders' monitoring incentives (e.g., Ashcraft and Santos 2009; Subrahmanyam et al. 2014; Martin and Roychowdhury 2015; Amiram et al. 2017), we predict that this effect is stronger for nonrelationship lead arrangers. Because non-relationship arrangers initiate new relationships with the borrowers only once CDSs become available, they are more likely to rely on CDSs, rather than on borrower monitoring, to protect themselves against credit risk. Relative to relationship arrangers, with their extensive knowledge of a borrower and deep channels of communication with its managers, monitoring new borrowers will be also substantially costlier for nonrelationship arrangers. Moreover, if non-relationship lead arrangers monitor less intensively, this may help them attract new borrowers, as borrowers prefer fewer constraints on their actions (e.g., Berlin and Mester 1992; Bradley and Roberts 2015). As non-relationship arrangers are likely to cross-sell additional services to these borrowers, allowing them greater flexibility to engage in new business activities would further benefit these lenders. Our fourth hypothesis is as follows.

H4: Following CDS initiation, non-relationship lead arrangers have weaker monitoring incentives, relative to relationship lenders.

2.5 Agency Problems within the Loan Syndicate

Syndicate participants typically maintain an arm's length relationship with a borrower and delegate information collection, contractual negotiations, and borrower screening and monitoring to the lead arranger. Therefore, participants face substantial agency issues when they join a syndicate (e.g., Lee and Mullineaux 2004; Sufi 2007; Ivashina 2009). Building on our fourth research hypothesis, we expect greater agency problems within the loan syndicate when it is arranged by a non-relationship lead arranger after CDS initiation. A stronger decline in the monitoring incentives of non-relationship relative to relationship lead arrangers should cause syndicate participants to be concerned that the former will shirk more on their monitoring duties (e.g., Sufi 2007; Ivashina 2009). These agency problems can be significantly mitigated if lenders with prior relationships with the borrower become syndicate participants, because they are more informed about the borrower's performance and creditworthiness (Sufi 2007). Thus, although we predict that non-relationship participants are more likely to join the syndicate after CDS initiation, we expect that this effect will be weaker when a loan is syndicated by a non-relationship lead arranger. Building on these arguments, we state our final hypothesis as follows.

H5: The increase in the probability that a non-relationship syndicate participant joins the loan syndicate following the initiation of CDS trading on a borrower's debt is lower if the loan is arranged by a non-relationship lead arranger.

3. Sample, Data, and Descriptive Statistics

3.1 Data sources and sample selection

We obtain syndicated loan characteristics from DealScan and CDS data from Markit. Borrowers' characteristics are from Compustat and CRSP, and banks' characteristics are from Call Reports (Form Y-9C regulatory filings). We collect firm credit ratings from Compustat and Mergent FISD, data on analyst coverage from I/B/E/S, data on media articles from RavenPack News Analytics, and data on mergers and acquisitions from SDC.

We start sample selection with all syndicated loans issued from 1994 to 2015, as reported in DealScan. There are 46,061 loan packages (deals) issued to US firms over this period. We match this sample to Compustat and eliminate observations with insufficient loan and firm data, resulting in 29,559 loan packages, issued to 6,956 firms. We match this sample to Markit and identify firms with traded CDSs at any point over our sample period (CDS firms hereafter). Following prior literature (e.g., Ashcraft and Santos 2009; Amiram et al. 2017), we identify the date of CDS trading initiation as the first date when a CDS quote for a firm is provided by Markit.² In line with Amiram et al. (2017), we eliminate firms for which the first CDS trade date falls in January 2001, the first month of Markit's coverage, because of the ambiguity of the exact date of initiation for these firms. We also require each CDS firm to have at least one loan issued in both the pre- and post-CDS initiation periods and non-CDS firms to have at least one loan issued starting in 2001, when CDS firms in our sample begin to have traded CDSs. Our final sample contains 843 CDS firms and 2,814 non-CDS firms, corresponding to 23,316 loan packages. This process yields 27,547 (177,936) arranger-loan (participant-loan) observations.

3.2 Descriptive statistics

Table 1 presents descriptive statistics of the main variables used in our test. The mean value of *No Relationship - Lead Arranger* indicates that 38% of lead arrangers in our sample do not have a prior relationship with the borrower. For lead arrangers that syndicate loans to CDS firms following CDS initiation, we define the *No Relationship - Lead Arranger* indicator variable

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² Our main findings are unchanged when we identify CDS initiation date as the first date when the quote for a five-year USD-denominated CDS contract is provided by Markit, following Ashcraft and Santos (2009). These contracts are the most commonly traded in the CDS market.

to be equal to 1 if the lead arranger has not syndicated a borrower's loans before the CDS initiation date and 0 otherwise. For lead arrangers that syndicate loans to CDS firms prior to CDS initiation or to non-CDS firms, we define *No Relationship - Lead Arranger* to be equal to 1 if the lead arranger has not syndicated a borrower's loans before the issuance date of the loan under consideration and 0 otherwise. The mean value of *No Relationship - Participant* indicates that 44% of syndicate participants in our sample do not have prior relationships with the borrower. For syndicate participants in loans to CDS firms following CDS initiation, we define the *No Relationship - Participant* indicator variable to be equal to 1 if the participant has not participants in loans to CDS firms prior to CDS initiation date and 0 otherwise. For participants in loans to CDS firms prior to CDS initiation or to non-CDS firms, we define *No Relationship - Participant* to be equal to 1 if the participant has not participated in a borrower's loans before the issuance date of the loan under consideration and 0 otherwise.

For both the lead arranger and participant samples, the firms are relatively large, as reflected by the mean values of *Assets*, defined as the natural logarithm of total assets (detailed variable definitions are reported in Appendix A). The average ratio of net income to total assets (*ROA*) is 0.03 for both samples, and 21% (16%) of sample observations relate to firms experiencing losses (*Loss*) for the lead arranger (participant) samples. The mean leverage ratio (*Leverage*), measured by the ratio of total liabilities to total assets, is 0.64 (0.66), and the average interest coverage ratio (*Interest coverage*), measured as EBIT divided by the interest expense, is 11.02 (9.65), while mean asset tangibility (*Tangibility*), measured as property, plant, and equipment divided by total assets, is 0.60 (0.59) for the lead arranger (participant) sample. In

³ We acknowledge that defining *No Relationship* based on the CDS initiation date for CDS firm and based on loan issuance date for non-CDS firms may raise a concern about the validity of our findings. In Section 4.2, we use difference-in-differences approach, where we assign a pseudo-CDS initiation date to the matched non-CDS firms using the initiation date of the paired CDS firms. This allows us to measure *No Relationship* over the same period for each pair of CDS and matched non-CDS firms.

addition, 58% (72%) of the lead arranger (participant) sample observations relate to borrowers that are rated (*Rated*) and 33% (42%) to borrowers with investment-grade ratings (*Investment Grade*). We also report statistics for loan characteristics that serve as controls. For both samples, the mean and median value of loan size (*Amount*), measured by natural logarithm of the loan amount, suggest that sample observations relate to loans that are relatively large (*Amount*), have an average maturity (*Maturity*) of approximately four years, and are subject to less than two financial covenants (*#Covenants*). A majority of observations are characterized by performance pricing provisions (*PP*), and 10% (11%) of them have a loan guarantor (*Guarantor*) for the lead arranger (participant) sample.

4. Empirical Results

4.1 The effect of CDS Trading on Non-Relationship Lending

To examine our first hypothesis that lenders are more likely to issue loans to a new borrower following the CDS trading initiation on the borrower's debt, we estimate the following logit model.

No Relationship =
$$\beta_0 + \beta_1 POST + \beta_2 Assets + \beta_3 ROA + \beta_4 Loss + \beta_5 Leverage$$

+ $\beta_6 Interest\ Coverage + \beta_7 Tangibility + \beta_8 Rated$
+ $\beta_9 Investment\ Grade + \beta_{10} Amount + \beta_{11} Maturity + \beta_{12} Guarantor$
+ $\gamma Fixed\ Effects + \varepsilon$, (1)

where *No Relationship* is one of the two indicator variables — *No Relationship-Lead Arranger* and *No Relationship-Participant* — defined as previously. Our main variable of interest is the *POST* indicator variable, which equals to 1 if the loan is issued after the CDS trading initiation date and 0 otherwise (this variable takes the value of 0 for all loans to non-CDS firms). If non-

relationship lending increases after CDS trading initiation on a borrower's debt, we expect a positive and significant coefficient on *POST*.

We control for firm and loan characteristics that can affect lenders' decisions to initiate new lending relationships, including a firm's size, profitability, the incidence of losses, interest coverage, tangibility and credit rating characteristics, as well as loan size, maturity and whether a loan has a guarantor. All firm characteristics are defined as previously and measured in the year preceding a loan's issuance. For loan packages that contain more than one loan (facility), we follow Ball et al. (2008) and Ivashina (2009) and use characteristics of the largest one. Our findings are unchanged when we control for the weighted-average loan characteristics of all loans in the package, where weights are based on loan (facility) size.⁴

To aid identification strategy, we include in Model (1) firm and year fixed effects, allowing us to implement a difference-in-differences research design, because CDS initiation dates are staggered across CDS firms. Firm fixed effects control for the time-invariant differences between CDS and non-CDS firms, while year fixed effects control for the time-varying factors common to all sample firms (e.g., Bertrand and Mullainathan 2003; Sapienza 2002; Valta 2012; Amiram et al. 2017). We acknowledge that, because firms may issue loans in months both before and after the initiation date during the CDS initiation year, employing year-month fixed effects is a better strategy to identify initiation effects. Due to concerns regarding a large number of fixed effects in nonlinear models (e.g., Maddalla 1987; Greene 2004), as an additional specification, we estimate Model 1 with a linear probability model, where we substitute year fixed effects with

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⁴ We do not control for the number of covenants and performance pricing provisions, as these characteristics are typically determined during the loan negotiation and therefore cannot affect lenders' choice of whether to start a new lending relationship. In any case, in untabulated analyses, we find that our results are robust to the inclusion of these variables.

⁵ In section 4.2 below, we propensity-score match CDS firm to non-CDS firm, which allows us to assign the CDS initiation date to each matched non-CDS firm. Our results are robust to this standard difference-in-difference approach.

year-month fixed effects. Moreover, as many of our subsequent analyses incorporate interaction terms, linear probability estimations also mitigate concerns regarding the interpretation of interaction terms in our nonlinear model estimations (e.g., Norton et al., 2004). To correct for within-firm correlation in the error term, we cluster standard errors at the firm level.

We present our findings in Panel A of Table 2. In columns 1 and 2, we report the estimation of Model (1) employing Logit and OLS models, respectively, for the lead arranger sample. We find a positive and significant coefficient on *POST* for both Logit and OLS specifications. Economically, based on the OLS specification, the probability that a non-relationship lead arranger syndicates the loan is 1.39 times higher following the CDS trading initiation (note that we measure economic significance based on the OLS specification for all tests). We find similar results for syndicate participants, as reported in columns 3 and 4. The probability that a participant with no relationship with the borrower joins the syndicate is 1.28 times higher once CDSs become available. These results are consistent with our primary prediction that lenders are more likely to initiate a new lending relationship after the inception of CDSs referencing the borrower's debt.⁶

With respect to controls, the negative and significant coefficient on *Leverage* suggests that higher leverage deters non-relationship lenders. The positive and significant coefficient on *Tangible* for the lead arranger sample implies that higher asset tangibility attracts non-relationship arrangers. When borrowers issue longer maturity loans and those with guarantors, we find that it is more likely for these loans to be arranged by non-relationship lenders.

One concern that might arise when using our difference-in-differences research design is whether the parallel trend assumption holds. To address it, we augment Model (1) with an

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⁶ In untabulated analyses, we find that the effect of CDS trading on non-relationship lending is more pronounced for lead arrangers than syndicate participants.

indicator variable *Last Loan*, which is equal to 1 for the last loan issued before the CDS initiation date and 0 otherwise. If there exists a pre-tend of non-relationship lending before CDS initiation, we expect to find a higher likelihood of non-relationship lending for the last loan before initiation (Heider and Ljungqvist 2015; Amiram et al. 2017). As we report in Panel B, the coefficients on *Last Loan* are insignificant and negative across all model specifications, indicating that the likelihood of non-relationship lending is not higher for the last loan before CDS initiation date relative to the loans preceding this loan. This evidence suggests that the parallel trend assumption holds in our tests.

To further verify the robustness of our findings, we repeat our analyses using a short-window sample that limits observations of CDS firms to three years before and after CDS initiation. These analyses mitigate concerns that factors other than CDS initiation could drive our results. We report these tests in Panel C. Consistent with our main findings, the coefficients on *POST* is positive and significant for all lead arranger and participant specifications.

We provide additional support for our inferences by examining how the effect of CDSs on non-relationship lending varies with a borrower's information opacity. We expect hedging through CDSs and private information revealed by CDS spreads to be more valuable for lenders when they lend to opaque firms. To measure a borrower's information opacity, we rely on its analyst coverage, which is helpful for lenders in assessing borrower creditworthiness (Güntay and Hackbarth 2010; Mansi et al., 2011). *Low Coverage* is an indicator variable equal to of 1 if the number of equity analyst following the borrower is below the sample median in a year of a loan's issuance and 0 otherwise. We augment model (1) with *Low Coverage* and the interaction term *POST* × *Low Coverage*. As we report in Appendix B, Table B1, the coefficient on *POST* × *Low Coverage* is positive and significant for the lead arranger specifications, suggesting that the

effect of CDSs on the likelihood of a non-relationship lead arranger syndicating a loan is more pronounced for opaque borrowers. Following CDS initiation, lead arrangers are 1.21 times more likely to initiate a new lending relationship with opaque borrowers relative to more transparent borrowers. However, the coefficient on *POST* × *Low Coverage* is positive but insignificant for the participant specifications. This evidence suggests that, because lead arrangers perform the primary screening and monitoring of the borrower, they are likely to incur higher adverse selection costs when initiating a new relationship with an opaque firm.

4.2 Addressing Endogeneity Concerns

Although Model (1) includes firm and time fixed effects, we next employ a PSM approach to further assuage endogeniety concerns associated with fundamental differences between CDS and non-CDS firms. Following Martin and Roychowdhury (2015), we employ the following CDS initiation model to construct a matched sample of CDS and non-CDS firms.

CDS Initiation =
$$\beta_0 + \beta_1 Rated + \beta_2 Investment \ Grade + \beta_3 Assets + \beta_4 Leverage$$

+ $\beta_5 Profit \ Margin + \beta_6 Return \ Volatility + \beta_7 MTB + \varepsilon$, (2)

where CDS Initiation is an indicator variable equal to 1 for the year of CDS initiation and 0 otherwise. Rated, Investment Grade, Assets, and Leverage are as previously defined. Profit Margin is net income divided by sales. Return Volatility is the standard deviation of the firms' monthly stock return over the year. MTB is the market value divided by the book value of equity. All variables are measured in the year before the year under consideration. We tabulate the

interaction terms remains the same as in non-linear model estimations that we tabulate.

⁷ Although our findings with respect to the coefficient on the $POST \times Low\ Coverage$ interaction term are consistent across Logit and OLS specifications, to further address the concern regarding the interpretation of interaction terms in non-linear models, we apply Buis' (2010) methodology. We continue to find positive and significant coefficient on POST × Low Coverage (untabulated). We perform this robustness test for all other analyses that include interaction terms in non-linear models. We find that the direction and economic significance of coefficients on

estimation results of the first stage PSM model in Appendix B, Table B2, Panel A. In line with prior studies (Martin and Roychowdhury 2015; Kim et al. 2017), we find that the probability of CDS initiation is higher when a firm is rated, has an investment grade rating, and is larger. This probability is lower when a firm is more profitable.

Using the propensity score ("predicted probability of CDS initiation") estimated by the first-stage PSM model (2), we construct the one-to-one matched PSM sample by selecting for each CDS firm a non-CDS firm having the closest propensity score. The matched sample enables us to employ a standard difference-in-difference research design, because we can assign a pseudo-CDS initiation date to the matched non-CDS firms using the initiation date of the paired CDS firms. We allow non-CDS firms to be matched to multiple CDS firms to improve covariate balancing and retain sample size (e.g., Saretto and Tookes 2013). Our final sample contains 719 CDS firms and 395 matched non-CDS firms (referring to 161,831 lender-loan observations). We tabulate the results of the covariate balancing analyses in Panel B of Table B2. These analyses suggest that the PSM matched samples are well balanced across the covariates for most CDS initiation determinants. The CDS firms are larger than matched non-CDS firms at the 10% significance level, but, economically, the difference in firm size is relatively small, given the absolute firm size of the matched CDS and non-CDS firms (mean assets for CDS (non-CDS) firms are USD 4,447 million (USD 3,944 million).

Panel A of Table 3 presents the estimation results based on the PSM sample. We augment model (1) with $POST \times CDS$ Traded interaction term, where CDS Traded is equal to 1 for CDS firms and 0 for non-CDS firms (the CDS Traded indicator variable is subsumed by firm fixed effects). The coefficient on $POST \times CDS$ Traded is a difference-in-differences estimator and is expected to be positive. Consistently, we find that positive and significant coefficients on POST

imes *CDS Traded* for both lead arranger and syndicate participant specifications, supporting our main hypothesis that non-relationship lenders are more likely to extend loans to new borrowers once CDSs become available. Economically, a lead arranger (participant) is 1.17 (1.13) times more likely to initiate non-relationship lending following CDS initiation. To further examine the robustness of our findings, we conduct PSM analysis for a short-window sample that limits observations to three years before and after CDS initiation date (pseudo-initiation date for non-CDS firms). The results presented in Panel B corroborate our prior findings: coefficients on $POST \times CDS$ Traded are significantly positive in all specifications.

Overall, the results in Table 3 support our primary findings of a significant effect of CDSs on non-relationship lending. In addition to addressing endogeneity concerns, these analyses also mitigate concerns associated with the measurement of non-relationship lending for CDS and non-CDS firms in our primary analyses. While for loans of CDS firm issued following CDS initiation we measure lending relationship relative to the CDS initiation date, because this date cannot be defined for non-CDS firms, we define non-relationship lending for these firms based on whether a lender had a relationship with the firm before the current loan's issuance date. By assigning a pseudo-CDS initiation date to non-CDS firms, a standard difference-in-differences approach allows us to measure non-relationship lending over the same period for each pair of CDS and matched non-CDS firms.

In additional analyses, we address a concern that the onset of CDS trading itself may be endogenous. Following prior literature, we use lenders' foreign exchange (FX) derivative position as an instrument for CDS trading (e.g., Saretto and Tookes 2013; Subrahmanyam et al. 2014; Shan et al. 2015; Amiram et al. 2017). Because banks that hedge one component of their portfolio are more likely to hedge others (e.g., Minton et al. 2009), we expect banks that hedge

foreign exchange risk to also hedge credit risk via CDSs. Further, lenders' FX derivatives position is a macro, rather than a firm-specific, hedge (e.g., Saretto and Tookes 2013). Thus, while related to a bank's general propensity to hedge, the decision to hedge foreign exchange should be exogenous to a decision to start a lending relationship with a particular borrower.

To construct the *Forex* instrument, we collect notional amounts of FX derivatives from Call Reports filed by bank holding companies. For banks with available data, *Forex* is defined as the ratio of notional amounts of FX derivatives to total assets. We augment CDS initiation model (2) with *Forex* and present, in Appendix B, Table B3, the results of the second-stage IV regression estimated simultaneously with the first-stage model (2). For both specifications, the coefficients on *Post* are positive and significant. Although we cannot be certain that the *Forex* instrument is truly exogenous, the results of IV tests confirm that lead arrangers and participants are more likely to enter into a new lending relationship when CDSs are available on the borrower's debt.⁸

4.3 Borrower Cross-Selling Potential

In this section, we examine whether non-relationship lenders are more likely to pursue borrowers with higher cross-selling potential once CDSs become available. These analyses will help illuminate whether CDS-protected lenders pursue commercial aspects of the lending relationship when they initiate a new relationship. Since lenders' expectations about borrowers' future business needs are not directly observable, we employ multiple measures of high cross-selling potential to test our prediction.

Our first measure of cross-selling potential is based on the media coverage of borrowers'

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⁸ Due to Stata computational limitations, we cannot estimate the second-stage IV regression simultaneously with the first-stage model (2) when the second stage is estimated as the non-liner model with firm and year fixed effects or an OLS model with firm and month-year fixed effects. Therefore, to perform IV tests, we estimate the second-stage model as OLS with firm and year fixed effects.

business activities, as media is an important information source to non-relationship lenders (Bushman et al. 2016). We consider media articles that cover topics such as "public offering," "merger," "acquisition," and "investment" to relate to cross-selling opportunities. We focus on these topics because lenders benefit from cross-selling security underwriting and M&A advisory services, and receive significant originating fees if they syndicate loans to finance future investments of the borrower (e.g., Drucker and Puri 2005; Yasuda 2005; Ivashina and Kovner 2011; Fang et al. 2013; Euromoney 2014). Our discussions with loan officers at several banks suggest that lenders continuously monitor information in the media to gather information about borrowers' projected business activities and recent events, as these events may also indicate future business opportunities. Therefore, we consider media articles to be related to a borrower's cross-selling potential if they cover both past and expected borrower activities.

For each borrower in our sample, we obtain media articles from Ravenpack, which covers all news disseminated via Dow Jones Newswires. We limit media data to full-size articles with a relevance score of 75 and above. The relevance score is assigned by RavenPack to indicate when the firm is strongly related to the underlying news story (the scores range from 0 (low relevance) to 100 (high relevance)). We identify articles as cross-selling related if their news type, as assigned by Ravenpack, is public offering, merger, acquisition, or investment. We define the *High CS Potential-Media* indicator variable as equal to 1 if media articles about a borrower cover cross-selling related topics within three years before a loan's issuance date and 0 otherwise. We then augment model (1) with *High CS Potential-Media* and the interaction term $POST \times High \ CS \ Potential-Media$. Panel A of Table 4 presents our findings (we restrict these analyses to the post-2000 period because of Ravenpack's data availability). Consistent with our prediction that following CDS initiation lead arrangers are more likely to enter into a new

lending relationships if a borrower has higher cross-selling potential, we find the positive and significant coefficient on $POST \times High\ CS\ Potential-Media$ for the arranger specifications. Economically, non-relationship arrangers are 1.12 times more likely to syndicate loans to new borrowers with media coverage of cross-selling opportunities relative to borrowers without such coverage. The coefficient on $POST \times High\ CS\ Potential-Media$ is insignificant in participant specifications. This finding is likely to be explained by lead arrangers being the primary beneficiaries of cross-selling opportunities (e.g., Ivashina and Kovner 2011), as they maintain the direct relationship with the borrower.

Because M&A is a major line of business for investment banks, which offers lucrative business opportunities, such as M&A advisory and related financial services (Liaw 2011), our second measure of cross-selling potential further addresses a borrower's M&A activity. Although our media-based cross-selling potential measure reflects publicly available information about these deals, lenders' may also have private information concerning a borrower's future M&A transactions. Because we cannot capture lender's private information, we focus on actual M&A deals that occurred after loan issuance.

We define the *High CS Potential-M&A* indicator variable to be equal to 1 if a borrower engages in M&A deals within three years following a loan's issuance date and 0 otherwise. We then augment model (1) with *High CS Potential-M&A* and the interaction term *POST* × *High CS Potential-M&A*. We report the results in Panel B of Table 5. Similar to our findings in Panel A, the coefficient on *POST* × *High CS Potential-M&A* is positive and significant for the lead arranger specifications. Economically, following CDS initiation, non-relationship lead arrangers are 1.23 times more likely to issue loans to new borrowers with M&A business opportunities relative to borrowers without such opportunities. We do not find that cross-selling potential

influences whether non-relationship syndicate participants join the syndicate.

Our third measure of high cross-selling potential is based on a borrower's industry, as certain industries offer unique business opportunities related to structured finance. In particular, the aviation, ship building, oil, refinery and gas, and telecom industries offer banks such opportunities, which have higher fees/margins and lower risk due to high asset collateralization (Oliver Wyman 2006). Lenders also enjoy additional cross-selling benefits by selling to borrowers in these industries related derivatives products, such as commodity derivatives (especially for oil, gas, and natural resources), interest rate derivatives, and foreign currency derivatives. We define the industry-based measure of cross-selling potential — the High CS Potential-Industry indicator variable — to be equal to 1 if a borrower is in one of the listed industries and 0 otherwise. As we report in Panel C of Table 4, the coefficient on the interaction term POST × High CS Potential-Industry is positive and significant for the lead arranger specifications, but is insignificant for the participant specifications (note that High CS Potential-*Industry* indicator is subsumed by firm fixed effects). Following CDS initiation, lead arrangers are 1.23 times more likely to start a new lending relationship with borrowers in industries with high cross-selling opportunities relative to borrowers in other industries. These findings reinforce our inference that, post CDS initiation, non-relationship lead arrangers are more likely to initiate lending relationship with borrowers that offer high cross-selling opportunities.

Our fourth measure of cross-selling potential is based on borrowers' use of derivativerelated products, which also provide lenders with lucrative cross-selling business (e.g., Oliver Wyman 2006; Euromoney 2014). Derivative transactions offer high margins, as lenders charge borrowers significant fees while entering into offsetting positions with other counterparties, which are typically less expensive than what borrowers are charged. Furthermore, derivative

⁹ We report in Appendix B, Table B4, four-digit SIC codes we used to classify these industries.

positions translate into lower risk-weighted assets, as actual exposures are much smaller than the notional amounts, allowing lenders to reserve less capital for these transactions (e.g., Neilson et al. 2017). To proxy for borrower future derivative needs, we presume that borrowers with high demand for derivatives products are likely to have more derivative-related words in their 10-K filings. We define the *High CS Potential-Derivatives* variable to be equal to 1 if the ratio of derivative-related word counts (e.g., "derivative," "hedging") to total words count in the 10-K filing in the year preceding a loan's issuance is greater (less) than the sample median and zero otherwise. We find that the coefficient on *POST* × *High CS Potential-Derivatives* is positive and significant for both lead arrangers and participant specifications. Post CDS initiation, lead arrangers (participants) are 1.14 (1.12) times more likely to initiate lending relationships with borrowers having higher demand for derivative products relative to other borrowers (Panel D).

Our final measure of cross-selling potential addresses the extent of a borrower's foreign operations. Foreign operations provide lenders with transaction banking opportunities (e.g., trade finance) as well as related foreign currency hedging demand from borrowers. We measure the extent of foreign operations by an indicator variable *High CS Potential-Foreign*, which is equal to 1 if the ratio of a number of a borrower's foreign operating segments to its total number of operating segments in the year preceding a loan's issuance is greater than the sample median and zero otherwise. As we report in Panel E, the coefficient on *POST* × *High CS Potential-Foreign* is positive and significant for all specifications. Following CDS initiation, lead arrangers (participants) are 1.15 (1.08) times more likely to initiate a new lending relationship with borrowers having more extensive foreign operations relative to other borrowers.

Overall, the results reported in Table 4 suggest that once CDSs become available lead arrangers and syndicate participants focus more on pursuing new borrowers with high cross-

selling potential. Our findings also imply that, while lead arrangers are likely to be the primary beneficiaries from cross-selling of security underwriting, M&A advisory, and structured finance products, participants may also enjoy some cross-selling opportunities, such as derivative and foreign operation transactions.

4.4 Lender Loan Growth and Profitability

We next examine our prediction that CDSs have a stronger effect on non-relationship lenders that experience low loan growth, as CDSs offer lenders more flexible credit risk management and alleviate capital constraints. We define an indicator variable *Low Loan Growth* to be equal to 1 if the lender's average loan growth over the three years before a loan's issuance is below the sample median and 0 otherwise. We augment model (1) with *Low Loan Growth* and the interaction term *POST* × *Low Loan Growth*. We present our findings in Panel A of Table 5 (these analyses are restricted for lenders with data available from call reports). Consistent with our predictions, the coefficient on *POST* × *Low Loan Growth* is positive and significant for both lead arranger and participant subsamples. CDS-protected arrangers (participants) are 1.23 (1.29) times more likely to initiate non-relationship lending when they experience lower loan growth.

We also expect lenders with low profitability to have stronger incentives to pursue new borrowers once CDS hedging become available, because they may benefit from both additional interest revenue and cross-selling fees. We define an indicator variable *Low Profitability* to be equal to 1 if the lender's average ROE over the three years before a loan's issuance is below the sample median and 0 otherwise. We augment model (1) with *Low Profitability* and the interaction term $POST \times Low Profitability$. As we report in Panel B of Table 5, the coefficient on $POST \times Low Profitability$ is positive and significant for all specifications, in line with our prediction. Following CDS initiation, lead arrangers (participants) are 1.52 (1.49) times more

likely to issue loans to new borrowers when they experience low profitability. Overall, the results presented in Table 5 suggest that lenders' low loan growth and profitability amplify the effect of CDSs on initiating lending relationships for both lead arrangers and syndicate participants.

4.5 Lenders' Monitoring Incentives

We next investigate how non-relationship lenders' incentives to initiate new relationships in the post CDS-initiation period affect their monitoring incentives. We focus here on the lead arranger-borrower relationships since lead arrangers monitor borrowers on behalf of syndicate participants. Because non-relationship arrangers are likely to protect their credit exposure via CDS hedging, while pursuing commercial aspects of new lending relationships, we predict that, following CDS initiation, they have lower monitoring incentives relative to relationship lenders.

4.5.1 Lead Arrangers' Control Right

Because lenders retain stronger control rights when they are planning to intensively monitor a borrower, following prior literature, we measure monitoring incentives by the strength of lenders' control rights (Roberts and Sufi 2009a, 2009b; Roberts 2015; Kim et al. 2017). Performance (income statement-based) covenants act as tripwires and are used primarily to allocate control rights to lenders when a borrower underperforms, while capital (balance sheet-based) covenants address agency conflicts by aligning ex-ante incentives of a borrower and lenders (Christensen and Nikolaev 2012; Christensen et al. 2016). Therefore, the higher (lower) proportion of performance covenants should be associated with lenders' stronger (weaker) control rights. We define *Performance Covenants* as the ratio of the number of performance covenants to the sum of performance and capital covenants in a loan contract.¹⁰

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¹⁰ In line with Christensen and Nikolaev (2012), we classify cash interest coverage ratio, debt service coverage ratio, level of EBITDA, fixed charge coverage ratio, interest coverage ratio, ratio of debt to EBITDA, and ratio of senior debt to EBITDA covenants as performance covenants, while quick ratio, current ratio, debt to equity ratio, loan to

Roberts and Sufi (2009a) also suggest that performance pricing provisions affect lenders' control rights, by influencing the allocation of bargaining power between a borrower and lenders in ex post renegotiations. Performance pricing provisions define a pricing grid, which links a loan's interest rate to a borrower's performance (Asquith et al. 2005), where the interest rate increases (decreases) when a borrower's performance deteriorates (improves). Because the interest rate increasing provision typically imposes a sharp increase in the interest rate when performance deteriorates, borrowers are incentivized to renegotiate the loan contract, thus allocating control rights to lenders. We define *Interest Increasing PP* to be equal to 1 if a loan contract contains the interest rate increasing performance pricing provisions and 0 otherwise. We estimate *Interest Increasing PP* based on the largest loan facility in the package.

We estimate the following model to examine the effect of CDS initiation on non-relationship lead arrangers' control rights. 11

Control Rights =
$$\beta_0 + \beta_1 POST + \beta_2 No$$
 Relationship-Lead Arranger
+ $\beta_3 POST \times No$ Relationship-Lead Arranger + $\beta_4 Assets + \beta_5 ROA$
+ $\beta_6 Loss + \beta_7 Leverage + \beta_8 Interest$ Coverage + $\beta_9 Tangibility$
+ $\beta_{10} Rated + \beta_{11} Investment$ Grade + $\beta_{12} Amount + \beta_{13} Maturity$
+ $\beta_{14} Guarantor + \beta_{15} PP + \beta_{16} \# Covenants + \gamma Fixed$ Effects + ε , (3)

where *Control Rights* is either *Performance Covenants* or *Interest Rate Increasing PP*, as defined above. *No Relationship-Lead Arranger* and all control variables are defined as previously. If non-relationship lead arrangers retain weaker control rights post CDS initiation, we expect a negative and significant coefficient on the interaction term $POST \times No$ Relationship-

value ratio, ratio of debt to tangible net worth, leverage ratio, senior leverage ratio, and net worth requirement covenants as capital covenants. Financial covenants are identical to all loan facilities in the package.

¹¹ If loan package has more than one lead arranger (6.5% of sample packages), all lead arrangers are accounted for in estimating model (3). Our results are unchanged when we exclude these deals from the analyses.

Lead Arranger. For covenant specifications, we estimate model 3 as Tobit (OLS) with firm and year (firm and year-month) fixed effects. For performance pricing specifications, we estimate model 3 as Logit (OLS) with firm and year (firms and year-month) fixed effects. ¹²

As we report in columns 1 and 2 of Panel A, Table 6, the coefficient on $POST \times No$ Relationship-Lead Arranger is negative and significant in both specifications for Performance Covenants, indicating that non-relationship lead arrangers impose a lower proportion of performance covenants following CDS initiation relative to relationship lenders. Economically, in the post-CDS initiation period, the proportion of performance covenants is 0.89 times lower when non-relationship lead arrangers syndicate a loan relative to when a loan is syndicated by relationship lead arrangers. We continue to find negative and significant coefficients on $POST \times$ No Relationship-Lead Arranger for Interest Rate Increasing PP specifications in column 3 and 4, suggesting that CDS-protected non-relationship lead arrangers are less likely to impose provisions that would increase the interest rate if a borrower underperforms. In the post-CDS initiation period, non-relationship lead arrangers are 0.87 times less likely to impose the interest rate increasing performance provisions relative to relationship lead arrangers. To support the robustness of these findings, in columns 5 and 6 of Panel A, we perform the analyses for the weighted average Interest Rate Increasing PP variable for all facilities in the deal, where weights are based on loan (facility) size. We find very similar results.

Note that, in contrast to the negative and significant coefficients on $POST \times No$ Relationship-Lead Arranger, the coefficients on Post, which captures the effect of CDSs on control rights of relationship lenders, are insignificant in all specifications. We infer that lead arrangers that had a prior relationship with the borrower do not relinquish their control rights

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¹² We exclude #Covenants (PP) from the model when we examine performance covenants (interest increasing provisions), as it is highly correlated with the dependent variable.

after CDS initiation. This evidence is also noteworthy, as it relates to prior studies showing that following CDS initiation lenders retain weaker control rights (e.g., Shan et al. 2015; Chakraborty et al. 2015). We show that this result is actually attributed to lenders that initiate new relationships following CDS initiation.

In untabulated analyses, we re-estimate model (3) with covenant slack as the dependent variable, as covenant tightness also determines lenders' control rights (e.g., Dichev and Skinner 2002). Because financial ratios used in covenant definitions are adjusted extensively by lenders (e.g., Dichev and Skinner 2002; Li 2016; Li et al. 2016), we focus on the debt-to-EBITDA covenant, which slack can be measured reliably based on the DealScan data (e.g., Demerjian and Owens 2016). This covenant is also the most frequently used for our sample loans. We measure covenant tightness as the difference between a borrower's debt-to-EBITDA ratio in the year preceding the loan issuance measure minus the covenant threshold, divided by standard deviation of the debt to EBITDA ratio over the previous 12 quarters (e.g., Dichev and Skinner 2002). We find that, following CDS initiations, non-relationship lenders impose looser debt-to-EBITDA covenants, as indicated by the positive and significant (albeit at the 10% level) coefficient on $POST \times No$ Relationship-Lead Arranger. This evidence further suggests that lenders retain weaker control rights when entering into new lending relationships following CDS initiation.

4.5.2 Lead Arrangers' Share

Prior studies show that having more skin in the game incentivizes the lead arranger to monitor a borrower more intensively (e.g., Sufi 2007; Ball et al. 2008; Ivashina 2009). Therefore, we supplement the analyses of lenders' control rights by investigating whether non-relationship lead arrangers retain a lower loan share relative to relationship lenders following CDS initiation.

Following prior studies (Ball et al. 2008; Ivashina 2009; Amiram et al. 2017), we perform

these analyses for the largest loan facility in the package and define *Lead Arranger Share* as the proportion of the loan retained by the arranger. We re-estimate model (3) with *Lead Arranger Share* as the dependent variables and report the results in column 1 of Panel B. Consistent with our predictions, we find negative and significant coefficients on *POST* × *No Relationship-Lead Arranger*. Economically, in the post-CDS initiation period, lead arranger share is 0.92 times lower when non-relationship lead arrangers syndicate a loan relative to when loan is syndicated by relationship lead arrangers. To support the robustness of these findings, in column 2 of Panel B, we perform the analyses for the weighted average *Lead Arranger Share* for all facilities in the deal, where weights are based on loan (facility) size. We find that our results are unchanged.

The analyses presented in Panel B also have interesting implications with respect to Amiram et al.'s (2017) finding that lead arrangers increase their loan share following CDS initiation to compensate syndicate participants for the reduction in their monitoring incentives. In contrast to the negative and significant coefficient on *POST* × *No Relationship-Lead Arranger*, the coefficient on *Post*, which reflects the effect of CDSs on the loan share of relationship lenders, is positive and significant. This evidence indicates that Amiram et al.'s findings are attributed primarily to the increase in the skin in the game of lenders who had an established relationship with the borrower before CDS initiation.

Overall, the findings presented in Table 6 suggest that non-relationship arrangers have lower monitoring incentives relative to relationship arrangers after CDS initiation, consistent with these lenders relying more on credit protection via CDSs and taking more commercial perspective when initiating lending relationships once CDSs become available.

4.6 Resolving Agency Problems within the Loan Syndicate

The stronger negative influence of CDSs on monitoring incentives of non-relationship lead

arrangers is expected to worsen agency problems between these arrangers and syndicate participants. Because agency problems are typically less severe when participants had a prior lending relationship with the borrower, we predict that the effect of CDSs on the syndicate participation of non-relationship lenders is weaker when the loan is syndicated by a non-relationship arranger. To test this prediction, we augment model (1) with *No Relationship-Lead Arranger* and the interaction term $POST \times No$ Relationship-Lead Arranger. If non-relationship participants are less likely to join the syndicate in the post-CDS initiation period when the loan is syndicated by non-relationship lead arrangers relative to when it is syndicated by relationship lead arrangers, we expect negative and significant coefficient on $POST \times No$ Relationship-Lead Arranger. We report supporting evidence in Table 7. Non-relationship participants are 0.85 times less likely to join syndicates arranged by non-relationship lead arrangers. Coupled with the positive and significant coefficient on Post, as in previous analyses, these findings indicate that, while non-relationship lenders are more likely to join the syndicate once CDSs become available, this effect is attenuated when a loan is syndicated by a non-relationship lead arranger.

Agency problems within the syndicate can also be mitigated if participants have a prior relationship with the lead arranger (e.g., Sufi 2007; Bushman et al. 2017). In untabulated analyses, we examine whether participants familiar with the lead arranger are more likely to join the syndicate when the lead arranger initiates a lending relationship following CDS initiation. Although we do not find supporting evidence, this result is generally consistent with Sufi (2007), who shows that, when agency problems are severe, previous lead arranger-participant relationships are much less important than the previous relationships between the borrower and the participant.

¹³ Note that the sum of the coefficients on POST and $POST \times No$ Relationship – Lead Arranger is positive and significant, suggesting that, even when a loan is syndicated by a non-relationship lead arranger, the probability of non-relationship participants joining the syndicate is higher in the post- relative to the pre-CDS initiation period.

4.7 Supplementary Analyses

We supplement our analyses with exploratory tests of the effect of CDSs on the interest rate and size of loans syndicated by non-relationship lead arrangers. On the one hand, these arrangers may bid aggressively for a loan by offering a low spread or more credit. On the other hand, hedging loan exposure via CDSs is costly. If non-relationship arrangers charge a low spread or have a high loan exposure, the new relationship may not be profitable enough to pursue.

We re-estimate model (3) with *Interest Spread* as the dependent variable and report the results in Table 8. *Interest Spread* is the interest rate spread on the largest loan facility in the package. For robustness, we also perform the analyses for the weighted average *Interest Spread* for all facilities in the deal, where weights are based on loan (facility) size. The coefficient on the interaction term $POST \times No$ *Relationship-Lead Arranger* is insignificant in the interest spread specifications, suggesting that non-relationship lenders do not charge a lower spread relative to relationship lenders following CDS initiation.

We next re-estimate model (3) with *Amount* as the dependent variables, where *Amount* is the natural logarithm of the loan amount of the largest facility in the loan deal. We also perform analyses for the *Deal Amount* variable, measured by the natural logarithm of loan deal (package) amount. Interestingly, the coefficient on *POST* × *No Relationship-Lead Arranger* is negative and significant in both loan amount specifications. Economically, in the post-CDS initiation period, the size of the largest loan (deal) size is 0.92 (0.93) times lower when a non-relationship lead arranger syndicates a loan relative to when it is syndicated by a relationship lead arranger. This evidence indicates that CDS-protected non-relationship lead arrangers have lower exposures to new borrowers, which allows them to reduce hedging costs via CDSs. Given our findings of a substantial influence of cross-selling potential on lead arrangers' decisions to enter new

relationships, we infer that, although these lenders have a lower exposure to new borrowers to reduce hedging costs, borrowers still provide them with opportunities to serve their additional financial needs.¹⁴

5. Conclusion

We examine how the development of the CDS market affects lenders' incentives to initiate new lending relationships. Due to the information advantage of incumbent relationship lenders, non-relationship lenders face adverse selection problems when competing for a loan deal, which inhibits their relationship with new borrowers (Rajan 1992). We predict that an opportunity to hedge against credit losses and revelation of private information through CDS spreads significantly mitigate non-relationship lenders' adverse selection concerns, increasing their willingness to compete for loans for new borrowers. Consistent with our prediction, we find that both lead arrangers and syndicate participants are more likely to initiate new relationships once CDSs become available. Our findings also highlight the importance of a borrower's cross-selling potential in initiating new lending relationship. CDS-protected non-relationship lenders pursue commercial aspects of lending relationships and are more likely to enter into a new relationship if a borrower has high cross-selling potential. Furthermore, we show that lenders with low loan growth or profitability are more likely to initiate new lending relationships, because CDSs offer more flexible credit risk management and new relationships can boost profitability by generating additional interest revenue and cross-selling fees. We further find that, relative to relationship lead arrangers, non-relationship lead arrangers have lower monitoring incentives following CDS initiation, consistent with them protecting their credit exposure via CDS hedging, while focusing on providing additional noncredit services to the borrower. Lastly, we show that the increase in

¹⁴ We acknowledge that lower deal size may result in lower loan origination fees for lead arrangers, but expect cross-selling fess to be substantially higher than the decrease in origination fees.

the likelihood of non-relationship participants joining the syndicate following CDS initiation is attenuated if a loan is syndicated by a non-relationship lead arranger, which mitigates agency problems within the syndicate exacerbated by the lower monitoring incentives of non-relationship lead arrangers.

Our paper contributes to the growing literature on the influence of the CDS market on private lending by exploring how CDSs affect the formation of new lending relationships and lending strategies of non-relationship lenders. We further contribute to the expansive literature on relationship lending by showing that CDS availability increases the willingness of non-relationship lenders to compete for new borrowers and thus enhances the competition in the loan market. By documenting that a borrower's cross-selling potential significantly affects lenders' decisions to enter new relationships, we also add to the growing literature on the importance of noncredit business opportunities in private lending.

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APPENDIX A Variable Definitions

Variable		Definition
Amount	=	The natural logarithm of loan amounts of the largest facility in the loan package (DealScan).
Assets	=	The natural logarithm of total assets, measured in the year preceding a loan's issuance (Compustat).
CDS Initiation	=	An indicator variable equal to 1 for the year of CDS initiation and 0 otherwise (Markit).
Deal Amount	=	The natural logarithm of the package amount (DealScan).
Forex	=	The ratio of notional amounts of FX derivatives to the lenders' total assets, measured in the year preceding a loan's issuance. (Call Report).
Guarantor	=	An indicator variable equal to 1 if the loan is guaranteed and 0 otherwise (DealScan).
High-CS Potential-Derivatives	=	An indicator variable equal to 1 if the ratio of derivative-related word count to total words count in the 10-K filing in the year preceding a loan's issuance is greater than the sample median and 0 otherwise (10-K).
High-CS Potential-Foreign	=	An indicator variable equal to 1 if the ratio of a number of a borrower's foreign operating segments to its total number of operating segments in the year preceding a loan's issuance is greater than the sample median and 0 otherwise (Compustat).
High-CS Potential-Industry	=	An indicator variable equal to 1 if a borrower is in aviation, ship building, oil, refinery, gas, and telecommunication industries and 0 otherwise (Compustat).
High-CS Potential-M&A	=	An indicator variable equal to 1 if a borrower engages in M&A deals within three year after a loan's issuance and 0 otherwise (SDC).
High-CS Potential-Media	=	An indicator variable equal to 1 if a media covers topics related to cross-selling opportunities (e.g., Public Offering, Merger, Acquisition and Investment) within three years before a loan's issuance, and 0 otherwise (RavenPack).
Interest Coverage	=	The ratio of earnings before interest and taxes to interest expense, measured in the year preceding a loan's issuance. (Compustat).
Interest Spread	=	The natural logarithm of the all-in-drawn spread of the largest facility in the package (DealScan).
Interest Spread (weighted average)	=	The natural logarithm of the weighted average all-in-drawn spread for all facilities in the deal, where weights are based on loan (facility) size (DealScan).
Interest Rate Increase PP	=	An indicator variable equal to 1 if a loan contains an interest rate increasing performance pricing provision and 0 otherwise (DealScan).
Investment Grade	=	An indicator variable equal to 1 if a firm has an investment grade rating in the year prior to a loan's issuance and 0 otherwise (Compustat, Mergent FISD).
Last Loan	=	An indicator variable equal to 1 for the last loan issued prior to a borrower's CDS initiation date and 0 otherwise (DealScan, Markit).
Lead Arranger Share	=	The proportion (%) of the loan retained by the lead arranger, measured for the largest facility in the loan package (DealScan).

APPENDIX A (continued)

Variable		Definition
Lead Arranger Share	=	The weighted average <i>Lead Arranger Share</i> variable for all facilities in the
(weighted average)		deal, where weights are based on loan (facility) size (DealScan).
Leverage	=	The ratio of total liabilities to total assets, measured in the year preceding a
		loan's issuance. (Compustat).
Loss	=	An indicator variable equal to 1 if a borrower's net income in the year
		preceding a loan's issuance is less than zero, and otherwise (Compustat).
Low Coverage	=	An indicator variable equal to 1 if the number of equity analyst following a
		borrower in the year of a loan's issuance is less than the sample median and
		0 otherwise (I/B/E/S).
Low Loan Growth	=	An indicator variable equal to 1 if the lender's average loan growth over the
		three-year period prior to a loan's issuance is below the sample median and
		0 otherwise (Call Report).
Low ROE	=	An indicator variable equal to 1 if the lender's average ROE over the three-
20,7 110 2		year period prior to a loan issuance is below the sample median and 0
		otherwise (Call Report).
Maturity	=	The maturity of a loan in months (DealScan).
MTB	=	The market value divided by the book value of equity, measured in a year
N D I I . I . I .		prior to the year under consideration (Compustat).
No Relationship-Lead Arranger	=	For lead arrangers that syndicate loans to CDS firms following CDS
		initiation: an indicator variable equal to 1 if the lead arranger has not
		syndicated a borrower's loans prior to the CDS initiation date and 0
		otherwise. For lead arrangers that syndicate loans to CDS firms prior to
		CDS initiation or to non-CDS firms: an indicator variable equal to 1 if the
		lead arranger has not syndicated a borrower's loans prior to the issuance
		date of the loan under consideration and 0 otherwise (DealScan, Markit).
No Relationship-Participant	=	For syndicate participants in loans to CDS firms following CDS initiation:
		an indicator variable equal to 1 if the participant has not participated in a
		borrower's loans prior to the CDS initiation date and 0 otherwise. For
		syndicate participants in loans to CDS firms prior to CDS initiation or to
		non-CDS firms: an indicator variable equal to 1 if the participant has not
		participated in a borrower's loans prior to the issuance date of the loan
		under consideration and 0 otherwise (DealScan, Markit).
:Performance Covenants	=	The ratio of the number of performance covenants divided the sum of
		performance and capital covenants in a loan contract (DealScan).
POST	=	An indicator variable equal to 1 if the loan is issued after the CDS trading
		initiation date and 0 otherwise (Markit).
PP	=	An indicator variable equal to 1 if the loan has a performance pricing
		provision and 0 otherwise (DealScan).
Profit Margin	=	The ratio of net income to sales, measured in a year prior to the year under
y o		consideration (Compustat).
Rated	=	An indicator variable equal to 1 if a firm is rated by S&P or Moody's in the
Rated	=	An indicator variable equal to 1 if a firm is rated by S&P or Moody's in the vear prior to a loan's issuance and 0 otherwise (Compustat, Mergent FISD).
Rated Return Volatility	=	An indicator variable equal to 1 if a firm is rated by S&P or Moody's in the year prior to a loan's issuance and 0 otherwise (Compustat, Mergent FISD). The standard deviation of a firm's monthly stock return, measured in a year

APPENDIX A (continued)

Variable		Definition
ROA	=	The ratio of net income to total assets (Compustat).
Tangibility	=	The ratio of property, plant and equipment to total assets (Compustat).
#Covenants	=	The total number of financial covenants (DealScan).

Appendix B Additional Analyses

TABLE B1 Borrower Information Opacity

This table examines whether the effect of CDS trading on non-relationship lending is more pronounced for more opaque borrowers. Column 1 and 2 (3 and 4) report the analysis for the lead arranger (participant) specifications. Column 1 and 3 (2 and 4) present results using Logit (OLS) model. We include firm and year (year-month) fixed effects in Logit (OLS) model. t-statistics in parenthesis are based on standard errors clustered by firm level. ***, **, * indicates significance at the 0.01, 0.05, 0.10 level respectively. All other variables are defined in Appendix A.

		ıtionship		ıtionship
	- Lead A	- Lead Arranger		icipant
	(1)	(2)	(3)	(4)
POST	0.534***	0.092***	0.499***	0.112***
	(3.86)	(3.57)	(6.84)	(7.09)
Low Coverage	-0.179**	-0.035**	-0.055	-0.012
	(-2.43)	(-2.33)	(-0.88)	(-0.94)
POST*Low Coverage	0.338**	0.079***	0.092	0.018
	(2.30)	(2.82)	(1.16)	(1.06)
Model	Logit	OLS	Logit	OLS
Controls	Yes	Yes	Yes	Yes
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	24,582	27,196	175,677	177,815
Adj. (Pseudo) R ²	0.029	0.181	0.015	0.145

TABLE B2 Propensity Score Matching

This table presents the propensity score matching (PSM) analyses. Panel A provides results of the first stage probit model of *CDS initiation*. Panel B reports the difference in the means of the explanatory variables between the CDS firms and matched non-CDS firms to provide evidence of covariate balancing in the PSM estimation. t-statistics in parenthesis are based on standard errors clustered by firm level. ***, **, * indicates significance at the 0.01, 0.05, 0.10 level respectively. All variables are defined in Appendix A.

Panel A: The First Stage CDS Initiation Probit Model

CDS Initiation	
Rated	0.848***
	(13.46)
Investment Grade	0.362***
	(6.39)
Assets	0.107***
	(5.44)
Leverage	-0.019
	(-0.20)
Profit Margin	-0.002***
	(-3.03)
Return Volatility	0.075
	(1.14)
MTB	0.004
	(0.96)
Model	Probit
Observations	27,553
Adj. (Pseudo) R ²	0.192

TABLE B2 (continued) Propensity Score Matching

Panel B: Covariate Balancing

		Means	Difference in means
	CDS firms	non-CDS firms	(t-stats)
Rated	0.910	0.901	0.009
			(0.54)
Investment Grade	0.565	0.587	(0.022)
			(-0.85)
Assets	8.407	8.283	0.124*
			(1.77)
Leverage	0.647	0.651	(0.004)
			(-0.41)
Profit Margin	(1.446)	(0.003)	(1.443)
			(-0.98)
Return Volatility	0.043	0.041	0.002
			(0.14)
MTB	3.020	3.090	(0.070)
			(-0.30)

TABLE B3 Instrumental Variable (IV) Analyses

This table presents results of the second stage instrumental variable (IV) regression analyses, estimated simultaneously with the first stage CDS initiation model. Columns 1 (2) reports the analysis for the lead arranger (participant) specification. t-statistics in parenthesis are based on standard errors clustered by firm level. ***, **, * indicates significance at the 0.01, 0.05, 0.10 level respectively. All variables are defined in Appendix A.

	No Relationship	No Relationship
	- Lead Arranger	- Participant
	(1)	(2)
POST	1.209**	10.114***
	(2.11)	(4.19)
Assets	-0.040	-0.393**
	(-1.62)	(-2.15)
ROA	-0.025	-0.366
	(-0.17)	(-0.41)
Loss	0.000	0.032
	(-0.01)	(0.20)
Leverage	-0.052	-1.276**
	(-0.65)	(-2.39)
nterest Coverage	0.000	0.006**
	(0.63)	(2.42)
Tangibility	0.223*	1.420**
	(1.77)	(2.43)
Rated	-0.023	0.132
	(-0.51)	(0.56)
nvestment Grade	-0.059	-0.497*
	(-1.31)	(-1.90)
Amount	-0.022**	0.034
	(-2.09)	(0.46)
Maturity	0.001	-0.006*
	(1.44)	(-1.91)
Guarantor	0.031	0.095
	(1.25)	(0.69)
Model	OLS	OLS
Firm FE	Yes	Yes
Year FE	Yes	Yes
Observations	Yes	Yes
Adj. (Pseudo) R ²	8,483	33,759

TABLE B4 High Cross-Selling Potential Industries

This table presents SIC codes of industries we classify as having high cross-selling potential.

Industry	SIC code	SIC Industry Name
	1311	CRUDE PETROLEUM & NATURAL GAS
	1381	DRILLING OIL & GAS WELLS
	1382	OIL & GAS FIELD EXPLORATION SERVICES
	1389	OIL & GAS FIELD SERVICES, NEC
Oil, Refinery and Gas	2911	PETROLEUM REFINING
	4922	NATURAL GAS TRANSMISSION
	4923	NATURAL GAS TRANSMISISON & DISTRIBUTION
	4924	NATURAL GAS DISTRIBUTION
	3720	AIRCRAFT & PARTS
	3721	AIRCRAFT
Aviation	3724	AIRCRAFT ENGINES & ENGINE PARTS
Aviauon	3728	AIRCRAFT PARTS & AUXILIARY EQUIPMENT, NEC
	4512	AIR TRANSPORTATION, SCHEDULED
	4522	AIR TRANSPORTATION, NONSCHEDULED
Telecommunication	4813	TELEPHONE COMMUNICATIONS
Ship building	3730	SHIP & BOAT BUILDING & REPAIRING

TABLE 1
Descriptive Statistics

This table provides descriptive statistics for the lead arranger and participant samples. All variables are defined in Appendix A.

	I	Lead Arranger Sample			Participant Sample			
	N	Mean	Median	SD	N	Mean	Median	SD
No Relationship	27,457	0.38	0.00	0.49	177,936	0.44	0.00	0.50
Assets	27,457	7.79	7.70	2.06	177,936	8.28	8.20	1.69
ROA	27,457	0.03	0.04	0.09	177,936	0.03	0.04	0.07
Loss	27,457	0.21	0.00	0.41	177,936	0.16	0.00	0.37
Leverage	27,457	0.64	0.63	0.23	177,936	0.66	0.65	0.21
Interest Coverage	27,457	11.02	3.55	26.81	177,936	9.65	3.79	22.36
Tangibility	27,457	0.60	0.56	0.40	177,936	0.59	0.55	0.40
Rated	27,457	0.58	0.00	0.49	177,936	0.72	1.00	0.45
Investment Grade	27,457	0.33	1.00	0.47	177,936	0.42	1.00	0.49
Amount	27,457	19.47	0.00	1.54	177,936	20.06	0.00	1.14
Maturity	27,457	46.69	19.52	25.65	177,936	48.33	20.03	23.40
Guarantor	27,457	0.10	55.00	0.30	177,936	0.11	60.00	0.31
PP	27,457	0.42	1.00	0.49	177,936	0.57	1.00	0.50
#Covenants	27,457	1.14	0.00	1.31	177,936	1.38	0.00	1.32

TABLE 2 The Effect of CDS trading on Non-Relationship Lending

This table examines the effect of CDSs on non-relationship lending. Panel A reports results of the primary tests. Panel B reports results of the analysis that investigates whether the parallel trend assumption holds for our tests. Panel C reports results for the short-window analysis, where observations of CDS firms are limited to 3 years before and after CDS initiation. Columns 1 and 2 (3 and 4) report the analysis for the lead arranger (participant) specifications. Columns 1 and 3 (2 and 4) present results using Logit (OLS) model. We include firm and year (year-month) fixed effects in the Logit (OLS) model. t-statistics in parenthesis are based on standard errors clustered by firm level. ***, **, * indicates significance at the 0.01, 0.05, 0.10 level respectively. All variables are defined in Appendix A.

Panel A: Primary tests

	No Rela	itionship	No Rela	tionship
	- Lead A	Arranger	- Part	icipant
	(1)	(2)	(3)	(4)
POST	0.778***	0.149***	0.563***	0.125***
	(8.12)	(8.46)	(10.79)	(11.39)
Assets	0.011	0.003	-0.180***	-0.037***
	(0.22)	(0.31)	(-4.98)	(-4.94)
ROA	-0.053	-0.002	0.049	0.007
	(-0.16)	(-0.03)	(0.17)	(0.11)
Loss	0.077	0.021	0.016	0.002
	(1.19)	(1.58)	(0.31)	(0.21)
Leverage	-0.263*	-0.054*	-0.206*	-0.039
	(-1.81)	(-1.87)	(-1.70)	(-1.50)
nterest Coverage	0.000	0.000	0.001*	0.000**
	(-0.24)	(-0.37)	(1.86)	(2.00)
Tangibility	0.505***	0.103***	0.019	0.009
	(3.02)	(3.18)	(0.17)	(0.38)
Rated	-0.014	0.000	0.032	0.006
	(-0.18)	(0.01)	(0.53)	(0.47)
nvestment Grade	0.023	0.007	-0.128**	-0.027**
	(0.24)	(0.39)	(-2.37)	(-2.37)
Mount	-0.157***	-0.033***	0.041*	0.008*
	(-5.42)	(-5.92)	(1.81)	(1.69)
A aturity	0.008***	0.002***	0.010***	0.002***
	(8.10)	(8.80)	(14.12)	(14.40)
Guarantor	0.230***	0.047***	0.243***	0.052***
	(3.29)	(3.40)	(5.12)	(5.19)
Model	Logit	OLS	Logit	OLS
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	24,582	27,196	175,677	177,815
Adj. (Pseudo) R ²	0.029	0.181	0.015	0.145

TABLE 2 (continued)
The Effect of CDS trading on Non-Relationship Lending

Panel B: Pre-Trend Analyses

		No Relationship		No Relationship		
	- Lead A	Arranger	- Participant			
	(1)	(2)	(3)	(4)		
Last Loan	-0.151	-0.029	-0.031	-0.002		
	(-1.18)	(-1.33)	(-0.47)	(-0.13)		
POST	0.727***	0.138***	0.551***	0.124***		
	(7.06)	(7.20)	(9.40)	(9.95)		
Model	Logit	OLS	Logit	OLS		
Controls	Yes	Yes	Yes	Yes		
Firm FE	Yes	Yes	Yes	Yes		
Year FE	Yes	No	Yes	No		
Year-Month FE	No	Yes	No	Yes		
Observations	24,582	27,196	175,677	177,815		
Adj. (Pseudo) R ²	0.029	0.181	0.015	0.145		

Panel C: Short Window Analyses

	No Rela	tionship	No Rela	tionship
	- Lead	- Lead Arranger		icipant
	(1)	(2)	(3)	(4)
POST	0.354***	0.073***	0.249***	0.050***
	(3.18)	(3.63)	(4.18)	(4.18)
Model	Logit	OLS	Logit	OLS
Controls	Yes	Yes	Yes	Yes
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	18,377	21,423	129,685	131,879
Adj. (Pseudo) R ²	0.025	0.145	0.018	0.161

TABLE 3 Propensity Score Matching

This table presents results of the analysis based on the propensity score matched sample. Panel A reports results for the full sample, while Panel B reports results for short window analyses, which limit observations to 3 years before and after the CDS initiation date (the pseudo CDS initiation date for non-CDS firms). Columns 1 and 2 (3 and 4) report the analysis for the lead arranger (participant) specifications. Columns 1 and 3 (2 and 4) present results using Logit (OLS) model. We include firm and year (year-month) fixed effects in the Logit (OLS) model. t-statistics in parenthesis are based on standard errors clustered by firm level. ***, **, * indicates significance at the 0.01, 0.05, 0.10 level respectively. All variables are defined in Appendix A.

Panel A: PSM Analyses - Full Sample

		No Relationship - Lead Arranger		tionship icipant
	(1)	(2)	(3)	(4)
POST	-0.102	-0.014	-0.221***	-0.042***
	(-0.85)	(-0.72)	(-3.27)	(-3.30)
POST*CDS Traded	0.367**	0.074***	0.266***	0.060***
	(2.13)	(2.64)	(2.83)	(3.36)
Model	Logit	OLS	Logit	OLS
Controls	Yes	Yes	Yes	Yes
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	16,366	17,284	143,972	144,533
Adj. (Pseudo) R ²	0.100	0.295	0.130	0.172

Panel B: PSM Analyses - Short Window

	No Rel	No Relationship - Lead Arranger		tionship
	- Lead			icipant
	(1)	(2)	(3)	(4)
POST	-0.069	-0.014	-0.205***	-0.034***
	(-0.47)	(-0.62)	(-3.00)	(-2.66)
POST*CDS Traded	0.335*	0.064**	0.203**	0.046**
	(1.66)	(1.99)	(2.07)	(2.38)
Model	Logit	OLS	Logit	OLS
Controls	Yes	Yes	Yes	Yes
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	6,476	7,875	68,340	68,954
Adj. (Pseudo) R ²	0.040	0.275	0.011	0.158

TABLE 4 Borrower Cross-Selling Potential

This table examines whether the effect of CDS trading on non-relationship lending is more pronounced when a borrower has a high cross-selling potential. Panels A, B, C, D, and E reports results of the analyses where a cross-selling potential is measured based on media coverage of public offering, merger, acquisition, and investment topics in borrower-specific articles preceding a loan's issuance (*High CS Potential-Media*), a borrower's M&A transactions following a loan's issuance (*High CS Potential-M&A*), a borrower's participation in aviation, ship building, oil, refinery, gas, and telecommunication industries (*High CS Potential-Industry*), a borrower's demand for derivative related products (*High CS Potential-Derivatives*), and the extent of a borrower's foreign operations (*High CS Potential-Foreign*), respectively. In all panels, Columns 1 and 2 (3 and 4) report the analysis for the lead arranger (participant) specifications. Columns 1 and 3 (2 and 4) present results using the Logit (OLS) model. We include firm and year (year-month) fixed effects in Logit (OLS) model. t-statistics in parenthesis are based on standard errors clustered by firm level. ***, **, * indicates significance at the 0.01, 0.05, 0.10 level respectively. All other variables are defined in Appendix A.

Panel A: Media Coverage

	No Relationship - Lead Arranger		No Relationship - Participant	
	(1)	(2)	(3)	(4)
POST	0.737***	0.135***	0.398***	0.088***
	(5.71)	(5.83)	(6.15)	(6.54)
High CS Potential-Media	-0.066	-0.010	0.089*	0.016*
	(-0.95)	(-0.79)	(1.99)	(1.74)
POST*High CS Potential-Media	0.250**	0.047**	-0.071	-0.010
	(2.03)	(2.12)	(-1.18)	(-0.81)
Model	Logit	OLS	Logit	OLS
Controls	Yes	Yes	Yes	Yes
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	19,025	22,148	141,898	143,912
Adj. (Pseudo) R2	0.027	0.219	0.012	0.151

TABLE 4 (continued) Borrower Cross-Selling Potential

Panel B: Mergers and Acquisitions

	No Relationship - Lead Arranger			itionship icipant
	(1)	(2)	(3)	(4)
POST	0.756***	0.145***	0.572***	0.127***
	(7.84)	(8.16)	(10.94)	(11.52)
High CS Potential-M&A	-0.152	-0.025	0.118	0.025
	(-1.38)	(-1.31)	(1.56)	(1.64)
POST* High CS Potential-M&A	0.489**	0.087**	-0.142	-0.034
	(2.13)	(2.10)	(-1.03)	(-1.13)
Model	Logit	OLS	Logit	OLS
Controls	Yes	Yes	Yes	Yes
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	24,582	27,196	175,677	177,815
Adj. (Pseudo) R ²	0.029	0.181	0.015	0.145

Panel C: Industry Types

	No Relationship		No Relationship	
	- Lead A	Arranger	- Participant	
	(1)	(2)	(3)	(4)
POST	0.706***	0.136***	0.538***	0.119***
	(7.10)	(7.34)	(9.83)	(10.42)
POST*High CS Potential-Industry	0.520**	0.089**	0.173	0.042
	(2.11)	(2.03)	(1.37)	(1.51)
Model	Logit	OLS	Logit	OLS
Controls	Yes	Yes	Yes	Yes
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	24,582	27,196	175,677	177,815
Adj. (Pseudo) R ²	0.029	0.181	0.015	0.145

TABLE 4 (continued) Borrower Cross-Selling Potential

Panel D: Derivative Transactions

	No Relationship		No Rela	tionship
	- Lead A	- Lead Arranger		icipant
	(1)	(2)	(3)	(4)
POST	0.597***	0.115***	0.395***	0.089***
	(4.81)	(4.94)	(5.73)	(6.02)
High CS Potential-Derivatives	-0.027	-0.006	-0.014	-0.004
	(-0.45)	(-0.50)	(-0.34)	(-0.43)
POST*High CS Potential-Derivatives	0.298**	0.055**	0.253***	0.054***
	(2.38)	(2.31)	(3.78)	(3.75)
Model	Logit	OLS	Logit	OLS
Controls	Yes	Yes	Yes	Yes
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	24,582	27,196	175,677	177,815
Adj. (Pseudo) R ²	0.029	0.181	0.015	0.145

Panel E: Foreign Operations

	No Relationship - Lead Arranger			itionship icipant
	(1)	(2)	(3)	(4)
POST	0.667***	0.128***	0.512***	0.114***
	(6.21)	(6.29)	(9.17)	(9.55)
High CS Potential-Foreign	-0.107	-0.025*	-0.045	-0.009
	(-1.46)	(-1.77)	(-0.91)	(-0.88)
POST*High CS Potential-Foreign	0.295**	0.056**	0.153**	0.034**
	(2.07)	(2.12)	(2.01)	(2.09)
Model	Logit	OLS	Logit	OLS
Controls	Yes	Yes	Yes	Yes
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	24,582	27,196	175,677	177,815
Adj. (Pseudo) R ²	0.029	0.181	0.015	0.145

TABLE 5
Lender Loan Growth and Profitability

This table examines whether the effect of CDS trading on non-relationship lending is more pronounced for lenders with low loan growth (*Low Loan Growth*) or low profitability (*Low Profitability*). Panel A and B report results of the analyses based on a lender's loan growth and profitability, respectively. Columns 1 and 2 (3 and 4) report the analyses for the lead arranger (participant) specifications. Column 1 and 3 (2 and 4) present results using Logit (OLS) model. We include firm and year (year-month) fixed effects in Logit (OLS) model. t-statistics in parenthesis are based on standard errors clustered by firm level. ***, **, * indicates significance at the 0.01, 0.05, 0.10 level respectively. All other variables are defined in Appendix A.

Panel A: Lender Loan Growth

	No Rela	tionship	No Relationship	
	- Lead Arranger		- Participant	
	(1)	(2)	(3)	(4)
POST	0.181	0.043	0.374***	0.066***
	(1.00)	(1.61)	(4.31)	(4.53)
Low Loan Growth	-0.408***	-0.062***	-0.362***	-0.062***
	(-4.32)	(-4.45)	(-9.71)	(-10.08)
POST*Low Loan Growth	0.382**	0.059**	0.464***	0.080***
	(2.03)	(2.23)	(6.79)	(6.67)
Model	Logit	OLS	Logit	OLS
Controls	Yes	Yes	Yes	Yes
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	6,149	9,333	34,909	36,589
Adj. (Pseudo) R ²	0.043	0.382	0.035	0.215

Panel B: Lender Profitability

	No Rela	No Relationship - Lead Arranger		itionship
	- Lead A			icipant
	(1)	(2)	(3)	(4)
POST	-0.139	0.001	0.170*	0.040***
	(-0.61)	(0.04)	(1.80)	(2.78)
Low Profitability	0.038	0.015	0.043	0.008
	(0.41)	(0.98)	(1.14)	(1.20)
POST*Low Profitability	0.912***	0.134***	0.812***	0.136***
	(3.57)	(3.66)	(9.42)	(9.52)
Model	Logit	OLS	Logit	OLS
Controls	Yes	Yes	Yes	Yes
Firm FE	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes
Observations	6,149	9,333	34,909	36,589
Adj. (Pseudo) R ²	0.044	0.176	0.038	0.150

TABLE 6 Non-Relationship Lenders' Monitoring Incentives

This table presents results of the analyses that examine non-relationship lead arrangers' monitoring incentives following CDS initiation. In Panel A, we measure monitoring incentives by the strength of lenders' control rights, where in columns (1) and (2) we examine the proportion of performance covenants in total of performance and capital covenants (*Performance Covenants*) and in columns (3) - (6) the existence of the interest rate increasing pricing provision (*Interest Rate Increasing PP*). In Panel B, we measure monitoring incentives by the share retained by the lead arranger (*Lead Arranger Share*). We include firm and year (year-month) fixed effects in Logit and Tobit (OLS) models. t-statistics in parenthesis are based on standard errors clustered by firm level. ***, **, * indicates significance at the 0.01, 0.05, 0.10 level respectively. All variables are defined in Appendix A.

TABLE 6 (continued)
Non-Relationship Lenders' Monitoring Incentives

Panel A: Lead Arrangers' Control Rights

aner A. Leau Arrangers		e Covenants	Interest Rate	Increasing PP		Increasing PP l average)
	(1)	(2)	(3)	(4)	(5)	(6)
POST	0.081*	-0.009	0.091	0.015	0.050	0.013
	(1.88)	(-0.65)	(0.82)	(0.89)	(0.45)	(0.81)
No Relationship	-0.004	-0.003	0.008	0.000	0.052	0.002
- Lead Arranger	(-0.16)	(-0.49)	(0.15)	(-0.07)	(0.99)	(0.33)
POST*No Relationship	-0.263***	-0.041***	-0.251**	-0.039**	-0.238**	-0.039**
- Lead Arranger	(-4.98)	(-2.98)	(-2.16)	(-2.37)	(-2.06)	(-2.42)
Assets	-0.280***	-0.014*	-0.135**	-0.015*	-0.114*	-0.014
	(-19.76)	(-1.94)	(-2.13)	(-1.72)	(-1.80)	(-1.64)
ROA	-0.077	-0.028	-0.767	-0.055	-0.581	-0.047
	(-0.46)	(-0.53)	(-1.53)	(-0.82)	(-1.15)	(-0.71)
Loss	0.163***	0.026**	-0.258***	-0.030**	-0.251***	-0.031***
	(4.67)	(2.30)	(-2.93)	(-2.45)	(-2.88)	(-2.59)
Leverage	0.361***	0.034	-1.502***	-0.173***	-1.474***	-0.172***
	(6.13)	(1.38)	(-7.01)	(-6.37)	(-6.98)	(-6.42)
Interest Coverage	0.002***	0.000*	-0.001	0.000	-0.001	0.000
	(5.17)	(1.88)	(-0.80)	(-0.75)	(-0.56)	(-0.62)
Tangibility	-0.351***	-0.007	0.030	-0.003	0.102	0.001
	(-8.88)	(-0.29)	(0.16)	(-0.10)	(0.54)	(0.04)
Rated	0.221***	0.008	0.092	0.020	0.044	0.018
	(6.48)	(0.66)	(0.84)	(1.28)	(0.41)	(1.21)
Investment Grade	-0.438***	-0.031**	0.415***	0.072***	0.403***	0.069***
	(-10.80)	(-2.04)	(3.99)	(4.09)	(3.88)	(4.02)
Amount	0.115***	0.022***	0.448***	0.056***	0.495***	0.056***
	(8.76)	(5.54)	(11.36)	(12.05)	(12.17)	(12.38)
Maturity	0.004***	0.001***	0.002	0.000	0.002	0.000
	(8.15)	(4.37)	(0.90)	(1.63)	(1.13)	(1.37)
Guarantor	0.681***	0.158***	0.382***	0.056***	0.414***	0.056***
	(19.47)	(11.64)	(4.35)	(4.01)	(4.74)	(4.13)
PP	1.388***	0.310***				
	(54.22)	(36.74)				
#Covenants			0.740***	0.109***	0.768***	0.109***
			(27.11)	(30.32)	(28.44)	(31.00)
Model	Tobit	OLS	Logit	OLS	Logit	OLS
Firm FE	Yes	Yes	Yes	Yes	Yes	Yes
Year FE	Yes	No	Yes	No	Yes	No
Year-Month FE	No	Yes	No	Yes	No	Yes
Observations	27,457	27,196	20,819	27,196	21,159	27,196
Adj. (Pseudo) R ²	N/A	0.517	0.164	0.298	0.180	0.307

TABLE 6 (continued) Non-Relationship Lenders' Monitoring Incentives

Panel B: Lead Arrangers' "skin in the game"

	Lead Arranger Share	Lead Arranger Share (weighted average)
	(1)	(2)
POST	3.449***	5.018***
	(3.63)	(5.26)
No Relationship	3.784***	2.871***
- Lead Arranger	(6.59)	(4.51)
POST*No Relationship	-3.334***	-3.532***
- Lead Arranger	(-2.96)	(-3.29)
Assets	-2.103***	-1.882**
	(-3.12)	(-2.58)
ROA	1.359	0.880
	(0.22)	(0.13)
Loss	3.246***	2.552**
	(3.47)	(2.49)
Leverage	0.169	1.122
	(0.06)	(0.40)
Interest Coverage	0.024*	0.033**
	(1.72)	(2.06)
Tangibility	-1.374	-1.880
	(-0.64)	(-0.86)
Rated	0.884	2.337*
	(0.71)	(1.86)
Investment Grade	0.862	1.977*
	(0.77)	(1.79)
Amount	-10.094***	-9.967***
	(-19.08)	(-16.10)
Maturity	-0.098***	-0.096***
	(-6.20)	(-5.76)
Guarantor	1.128	0.760
	(1.28)	(0.84)
PP	-6.405***	-5.318***
	(-8.29)	(-6.35)
#Covenants	-0.230	-0.841**
	(-0.67)	(-2.22)
Model	OLS	OLS
Firm FE	Yes	Yes
Year FE	No	No
Year-Month FE	Yes	Yes
Observations	8,398	8,398
Adj. (Pseudo) R^2	0.718	0.659

TABLE 7 Agency Problems within the Syndicate

This table presents results of the analyses that investigate how agency problems within the syndicate are mitigated following CDS initiation. Column 1 (2) present results using Logit (OLS) model. We include firm and year (yearmonth) fixed effects in Logit (OLS) model. t-statistics in parenthesis are based on standard errors clustered by firm level. ***, **, * indicates significance at the 0.01, 0.05, 0.10 level respectively. All other variables are defined in Appendix A.

	No Relationship - Participant		
	(1)	(2)	
POST	0.559***	0.120***	
	(10.12)	(10.46)	
No Relationship - Lead Arranger	0.745***	0.161***	
	(21.24)	(21.42)	
POST*No Relationship - Lead Arranger	-0.343***	-0.068***	
	(-5.57)	(-5.07)	
Model	Logit	OLS	
Controls	Yes	Yes	
Firm FE	Yes	Yes	
Year FE	Yes	No	
Year-Month FE	No	Yes	
Observations	175,677	177,815	
Adj. (Pseudo) R ²	0.025	0.154	

TABLE 8 Supplementary Analyses

This table presents results of the analyses that examine the effect of CDSs on the interest spread and size of loans syndicated by non-relationship lead arrangers. In Column (1) and (2), we examine the interest spread (*Interest Spread*) and in column (3) and (4), we examine the size of a loan (*Amount*, *Deal Amount*). Analyses include firm and year-month fixed effects. t-statistics in parenthesis are based on standard errors clustered by firm level. ***, **, * indicates significance at the 0.01, 0.05, 0.10 level respectively. All variables are defined in Appendix 1.

	Loan Spread (1)	Loan Spread (weighted average) (2)	Amount (3)	Deal Amount (4)
POST	0.040	0.038	-0.030	-0.043
	(1.55)	(1.49)	(-0.92)	(-1.26)
No Relationship	0.059***	0.058***	-0.058***	-0.038**
- Lead Arranger	(5.81)	(5.75)	(-3.84)	(-2.41)
POST*No Relationship	-0.020	-0.019	-0.084**	-0.077*
- Lead Arranger	(-0.81)	(-0.79)	(-2.19)	(-1.95)
Assets	-0.111***	-0.112***	0.462***	0.479***
	(-8.72)	(-8.76)	(18.47)	(18.95)
ROA	-0.468***	-0.454***	0.408**	0.441***
	(-4.96)	(-4.97)	(2.49)	(2.59)
Loss	0.164***	0.164***	-0.077***	-0.070**
	(8.02)	(8.03)	(-2.66)	(-2.36)
Leverage	0.313***	0.316***	0.020	0.081
	(8.48)	(8.50)	(0.28)	(1.10)
Interest Coverage	-0.001***	-0.001***	0.001***	0.001***
	(-6.83)	(-6.96)	(2.61)	(2.89)
Tangibility	-0.309***	-0.311***	-0.015	-0.003
	(-6.89)	(-7.09)	(-0.20)	(-0.03)
Rated	0.040*	0.037*	-0.100***	-0.109***
	(1.77)	(1.67)	(-3.61)	(-3.75)
Investment Grade	-0.203***	-0.194***	0.029	-0.001
	(-6.00)	(-6.20)	(0.76)	(-0.03)
Amount	-0.060***	-0.057***		
	(-6.46)	(-6.16)		
Maturity	0.000	0.000	0.004***	0.005***
	(1.19)	(1.28)	(6.90)	(8.73)
Guarantor	0.013	0.017	-0.066**	-0.072**
	(0.79)	(0.99)	(-2.36)	(-2.40)

Table 8 (continued)

Table 8 (Continued)				
PP	-0.102***	-0.102***	0.214***	0.197***
	(-7.10)	(-7.65)	(11.40)	(10.38)
#Covenants	0.055***	0.055***	0.031***	0.059***
	(11.09)	(11.25)	(4.34)	(7.75)
Model	OLS	OLS	OLS	OLS
Firm FE	Yes	Yes	Yes	Yes
Year FE	No	No	No	No
Year-Month FE	Yes	Yes	Yes	Yes
Observations	27,196	27,196	27,196	27,196
Adj. (Pseudo)) R ²	0.756	0.76	0.790	0.774